REVIEW PAPER

UROLITHIASIS

# Recent scoring systems predicting stone-free status after retrograde intrarenal surgery; a systematic review and meta-analysis

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Oktay Özman ORCID-ID: 0000-0003-2499-8947 Gaziosmanpasa Training and Research Hospital Urology Clinic Karayolları Neigh. 34000 Istanbul, Turkey phone: +90 554 501 2312 ozmanoktay@hotmail.com **Introduction** Several scoring systems and nomograms have been developed to predict the success of retrograde intrarenal surgery. But no meta-analysis for the performance of scoring systems has yet been performed. The aim of this study was to compare predictive ability of recent scoring systems for stone-free rate of retrograde intrarenal surgery.

**Material and methods** PubMed and Web of Science databases were searched systematically between April and May 2021. The scoring systems which were validated externally or studied at least by two different researcher groups were selected for further analysis. Of 59 records, 14 studies met the inclusion criteria (n = 4137). Area under curve (AUC) values of selected scoring systems were pooled in random or fixed effects. The I<sup>2</sup> test was used to quantify heterogeneity.

**Results** Eight, 5, 8, 4 and 3 studies included in meta-analyses for the modified Seoul National University Renal Stone Complexity Score (S-ReSC), R.I.R.S., Resorlu-Unsal Score (RUS), S.T.O.N.E., and Ito's Nomogram, respectively. We found pooled AUC values 0.709 (95% CI 0.670–0.748), 0.704 (95% CI 0.668–0.739), 0.669 (95% CI 0.646 to 0.692), and 0.771 (95% CI 0.724 to 0.818), for first four of them, respectively. Heterogeneity was very high to pool AUC values for Ito's nomogram.

**Conclusions** Although S.T.O.N.E. score showed higer pooled AUC value, this systematic review and meta-analysis has not revealed superiority of any scoring system. High heterogeneity between studies and dependencies between scoring systems make it difficult to design a comparative statistical model to generalize the findings. Also, limitations aside, neither scoring system has demonstrated good predictive/discriminative performance.

Key Words: kidney stone () retrograde intrarenal surgery () scoring system () nomogram () stone

# INTRODUCTION

Advances in surgical techniques and urological devices have made minimal invasive surgery for kidney stones more preferred in recent years. Shock wave lithotripsy, retrograde intrarenal surgery (RIRS), and percutaneous nephrolithotomy (PNL) have replaced open kidney stone surgery in many cases [1]. Which treatment should be chosen for which stone in the kidney is of critical importance. The key is to use the right weapon at the right time. Retrograde intrarenal surgery stands out with its low complication rate and high stone-free rate (SFR), especially for stones up to 2 cm [2]. It is known that the success rate of RIRS depends on multiple factors, such as stone burden, number, localization, renal calyceal anatomy [3].

Several scoring systems and nomograms have been developed with these factors to predict the success of RIRS [3]. Emerging data suggest that scoring systems might provide a preoperative prediction for the outcome of RIRS but none of them gained popularity and were not widely used. Similar findings have been reported from different authors for several scoring systems [4, 5]. To date, no meta-analysis for the performance of scoring systems has yet been performed. We therefore aimed to compare predictive ability of the scoring systems for SFR of RIRS.

# MATERIAL AND METHODS

## Search strategy and selection criteria of studies and scoring systems

This systematic review and meta-analysis was conducted following the Preferred Reporting Items for Systematic Review and Meta-analysis (PRISMA) guidelines [6]. Study protocol was registered with PROSPERO (number: CRD42021252873). PubMed and Web of Science databases were searched systematically between April and May 2021. The search strategy was developed with keywords in accordance with Population Intervention Comparison Outcomes (PICO) strategy. All keywords and details of the search strategy can be found in Table 1–2. For each database, the search strategy, removal of duplicates, title-abstract screening were performed by 2 different authors independently (OÖ and ÖC). After the resolution of discrepancies by head-to-head meeting, the remaining articles were evaluated by OÖ based on full text. Full text versions of studies were requested from the authors if needed. Finally, reference lists of all full texts were reviewed for further relevant studies.

The scoring systems which were validated externally or studied at least by two different researcher groups were selected for further analysis. First research strategy was validated by second research strategy including names of selected scoring systems in 'Comparison' components instead of general relevant keywords. But no additional articles were found. Only original studies which have area under curve (AUC) metadata of ROC curve analysis and written in English were included in the meta-analysis. Studies including patients who underwent bilateral RIRS or had a different simultaneous surgical procedure (for kidney stone or other indication) in the same session, and patients who underwent RIRS to remove encrusted ureteral stent or other foreign bodies were excluded. Also studies including only specific patient groups (peadiatric, elderly etc.) or only patients with renal abnormalities (horseshoe kidney, solitary kidney, transplanted kidney etc.) were excluded.

## **Quality assessment**

Included studies were assessed by two reviewers (HC and OÖ) independently. The risk of bias assessments were done according to modified version of the Quality In Prognostic Studies (QUIPS) tool's six criteria; study participation, study attrition, prognostic factor measurement, outcome measurement, confounding, and statistical analysis and reporting [7]. Final decisions about discrepancies were made after a head-to-head discussion between the scorers. Surgical outcome evaluations by kidney, ureter, and bladder (KUB) imaging, ultrasonography (US), computed tomography (CT), or second-look flexible ureterorenoscopy were considered to be appropriate. Because of consensus lack on the exact diameter of significant residual fragments in the literature, any cut-off value was not took account during quality assessment of reported outcomes.

Table 1.	First research	strategy befo	ore selection	of candidate
scoring	systems for rev	view		

Р		I		С		0
kidney stone OR	AND	retrograde intrarenal	AND	scoring system OR	AND	stone free OR
kidney calculi OR renal stone OR renal calculi OR urolithiasis		OR flexible ureteroreno- scopy OR flexible ureteroscopy		nomogram		residual stone
OR nephrolithiasis						
OR						
upper ureter stone						
OR						
upper ureter calculi						
OR						
proximal ureter stone						
OR						
proximal ureter calculi						

# Table 2. Second research strategy after selection of candidate scoring systems for review

Р		I		С		0
kidney stone OR kidney calculi OR	AND	retrograde intrarenal surgery OR flexible	AND	Resorlu-Unsal Score OR RUS	AND	stone free OR residual stone
renal stone OR renal calculi OR urolithiasis		ureteroreno- scopy OR flexible ureteroscopy		modified Seoul National University Renal Stone Complexity Score		
OR nephrolithiasis OR upper ureter stone				OR S-ReSC OR R.I.R.S		
OR upper ureter calculi OR				OR RIRS OR S.T.O.N.F.		
proximal ureter stone OR proximal				OR STONE OR Ito's nomogram		
ureter calculi						

PICO – Population Intervention Comparison Outcomes; RUS – Resorlu-Unsal Score; S-ReSC – Seoul National University Renal Stone Complexity

# Data extraction and analysis

Area under the ROC curve were pooled in random or fixed effects for five scoring systems separately. The I<sup>2</sup> test was used to quantify heterogeneity. A fixed-effects model was used where the I<sup>2</sup> was below 30%; otherwise, a random-effects model was used [8]. Area under curve values were pooled in MedCalc Statistical Software version 20 (MedCalc Software by, Ostend, Belgium). An AUC of  $\geq$ 0.800 was considered to denote reasonable/favourable discriminating ability. p values under 0.05 were considered statistically significant.

# RESULTS

Study selection flow chart, derived from the PRIS-MA 2020 Flow Diagram can be seen in Figure 1. The search yielded 83 articles. Following removal of 24 duplicate publications, 38 articles were excluded during title and abstract screening. To selection criteria of scoring systems, Resorlu-Unsal Score (RUS), modified Seoul National University Renal Stone Complexity Score (S-ReSC), R.I.R.S., S.T.O.N.E., and Ito's nomogram were selected and



Figure 1. Study selection flow chart.

n – number of patients; RIRS – retrograde intrarenal surgery; AUC – area under curve; ROC – receiver operating characteristic

#### Table 3. Summary of parameters included in scoring systems

Parameters	S-ReSC	RUS	R.I.R.S.	S.T.O.N.E	Ito's
Stone burden		+	+	+	+
Stone localization	+			+	
Number of stones		+		+	+
Stone in lower calyx	+	+			+
Operator experience					+
Hydronephrosis				+	+
Hounsfield Unit (HU)			+	+	
Infundibulopelvic angle (IPA)		+	+		
Infundibulopelvic length (IL)			+		
Abnormal renal anatomy		+			

S-ReSC – Seoul National University Renal Stone Complexity; RUS – Resorlu-Unsal Score

relevant studies included in the systematic review. (T)allness, (O)ccupied lesion and (HO)unsfield units (T.O.HO) score and 2 nomograms were excluded due to the lack of external validation [9, 10, 11]. Also one nomogram which developed for pediatric patients and one nomogram predicting perioperative complications were excluded [12, 13]. After exclusion of 7 article in the full-text evaluation process, remaining 14 studies were included in systematic review. Thirteen studies providing meta-data for quantitative analyses were included in one or more meta-analyses.

The characteristics of the selected studies were shown in Table 4. There were 3, 1, and 4 articles which studied 4, 3, and 2 different scoring systems at the same time. Five articles had evaluated only single scoring system. All scoring systems had been included in 3 or more studies. Most studied scoring systems were RUS and S-ReSC scores (8 times for each). Parameters which included in scoring

#### Table 4. Characteristics of studies included in systematic review

First author/		During	Stone	Number	Outcome	Postoperative	Stone-free status	S	tudied	scoring	g systems	5
year	Country	Design	stone burden	of patients	(SFR)	imaging me- thod	definition	S-ReSC	RUS	R.I.R.S.	S.T.O.N.E	Ito's
Selmi 2020	Turkey	Ρ	Kidney 924 mm³	110	(81/110) 73.6%	Not indicated	Not having residual stone fragments greater than 4 mm	+	+	+	+	
Bozkurt 2021	Turkey	R	Kidney 103 mm²	949	(743/949) 78.3%	СТ	Residual fragments <2 mm	+	+	+		+
Richard 2020	France	R	Kidney and/or upper ureter 11 mm	800	(593/800) 74.1%	Radioscopic imaging or CT	Total absence of residual stone	+	+		+	+
Ozbek 2020	Turkey	R	Kidney 13 mm	280	(215/280) 76.7%	СТ	Complete clearence	+	+	+		
Erbin 2016	Turkey	R	Kidney 145 mm²	339	(238/339) 70.1%	KUB, US or CT	No evidence of residual stones or fragments at 1 month follow-up	+	+			
Karsiyakali 2020	Turkey	R	Kidney and/or upper ureter 140 mm²	81	(60/81) 74.1%	KUB or CT	Clinically insignificant residual stones <4 mm	+			+	
Jung 2014	S. Korea	R	Kidney 12 mm	88	(75/88) 85.2%	СТ	No evidence of residual stone on post-operative CT for 1 month	+				
Park 2015	S. Korea	R	Kidney 1.6/2.5 cm³	159	(116/159) 73%	СТ	No evidence of a stone or with clinically insignificant residual fragments less than 2 mm	+				
Xiao 2017	China	R	Kidney 14 mm	382	(281/382) 73.6%	KUB or CT (if KUB showed any high- densities or radiolucent stones)	No detectable stone on KUB, and fragments of less than 2 mm		+	+		
Wang 2021	China	R	Kidney 13 mm	147	105/147 (71.4%)	KUB or CT	No detectable stone on KUB or non-contrast CT, or fragments of less than 2 mm		+	+		
Sfoungaristos 2016	Israel	Ρ	Kidney 10 mm	85	63/85 (74.1%)	СТ	The absence of any residual fragment		+			
Molina 2014	USA	R	Kidney and/or ureter 9 mm	200	164/200 (82%)	Intraoperative endoscopic inspection with fluoroscopy and/or CT	The absence of stone fragments or fragments ≤ 2 mm				+	
lto 2014	Japan	R	Kidney 679/3035 mm <sup>3</sup>	310	185/310 (59.7%)	СТ	The strict absence of visible stones on imaging					+
Resorlu 2012	Turkey	R	Kidney 16 mm	207	178/207 (86%)	Intraoperative endoscopic inspection, US, CT	CIRF ≤1 mm		+			

P – prospective; R – retrospective; CT – computed tomography; KUB – kidney ureter bladder; SRF – stone-free rate; CIRF – clinically insignificant residual fragment; US – ultrasonography; RUS – Resorlu-Unsal Score; SReSC – Seoul National University Renal Stone Complexity

systems were summarized in Table 3. These can be classified in three categories; stone related parameters (stone burden, localization, density etc.), anatomic parameters (infundibular measurements, hydronephrosis, abnormal anatomy etc.), and surgeon related parameter (experience).

## **Quality assessment**

The risks of bias assessments were displayed in Table 5. Prognostic factor and outcome measurements were the most poorly rated domains among six criteria of QUIPS tool. Except two of them, retrospective design of included studies caused a scoring system calculation bias risk. There weren't clear statements indicating exact calculation time of scoring systems in validation studies (preoperatively or in the retrospective study period). More importantly, most of the studies didn't report inter-/intraobserver variability analyses. Thus, most of the studies showed poor quality in terms of prognostic factor measurement criteria. Another common lack that reduces the quality of the studies was the non-standard outcome measurement.

## **Data analysis**

Relevant metadata which extracted from selected studies for each scoring system and results of metaanalyses with pooled data can be seen in Table 6. All scoring systems' AUC values showed statistical-

### Table 5. Risk of bias rating

ly significant heterogeneity in first inclusive metaanalyses (S-ReSC and RUS; moderate, R.I.R.S. and S.T.O.N.E.; high and Ito's nomogram; very high). **S-ReSC score.** Eight studies reported AUCs for S-ReSC score [4, 5, 15, 20–24]. Meta-analysis yielded a pooled AUC of 0.716 (95% CI 0.669 to 0.762). Het-

erogeneity was moderate ( $I^2 = 74\%$ ). It reduced to  $I^2 = 51\%$  (p = 0.09) after exclusion of three studies by Jung, Erbin, and Ozbek with new pooled AUC of 0.709 (95% CI 0.670 to 0.748, random effects) [15, 20, 21].

**R.I.R.S. score.** Five studies reported UACs for R.I.R.S. score [4, 5, 16, 20, 25]. Meta-analysis yielded a pooled AUC of 0.781 (95% CI 0.711 to 0.851). Heterogeneity was high ( $I^2 = 89\%$ ). It reduced to  $I^2 = 0\%$  (p = 0.43) after exclusion of Xiao's and Ozbek's studies with new pooled AUC of 0.704 (95% CI 0.668 to 0.739, fixed effects) [16, 20].

**RUS score.** Eight studies reported AUCs for RUS score [4, 5, 16, 20, 21, 22, 25, 26]. Meta-analysis yielded a pooled AUC of 0.711 (95% CI 0.668 to 0.754). Heterogeneity was moderate ( $I^2 = 75\%$ ). It reduced to  $I^2 = 6\%$  (p = 0.4) after exclusion of two studies by Xiao and Ozbek with new pooled AUC of 0.669 (95% CI 0.646 to 0.692, fixed effects) [16, 20].

**S.T.O.N.E. score.** Four studies reported AUCs for S.T.O.N.E [5, 18, 20, 22]. Meta-analysis yielded a pooled AUC of 0.728 (95% CI 0.647 to 0.809). Heterogeneity was high ( $I^2 = 88\%$ ). It reduced to  $I^2 = 22\%$  (p = 0.3) after exclusion of Richard's study with new pooled AUC of 0.771 (95% CI 0.724 to 0.818, fixed effects) [22].

Study	Study participation	Study attrition	Prognostic Factor Measurement	Outcome Measurement	Study Confounding	Statistical Analysis and Reporting
Selmi 2020	+	+	+	-	+	+
Bozkurt 2021	-	+	?	-	?	+
Richard 2020	+	?	?	+	?	?
Ozbek 2020	+	+	?	+	+	+
Erbin 2016	+	+	?	-	+	+
Karsiyakali 2020	+	+	-	-	?	+
Jung 2014	+	+	?	+	?	+
Park 2015	?	+	+	+	+	+
Xiao 2017	+	+	-	?	+	+
Wang 2021	+	+	-	-	+	+
Sfoungaristos 2016	+	+	-	+	+	+
Molina 2014	-	-	?	-	?	+
lto 2014	+	+	?	+	+	+
Resorlu 2012	?	+	?	+	+	?

Key: +; low risk bias,-; high risk bias, ?; unclear risk of bias

Studies including S-ReSC	AUC	SE	95% CI	Weight (%, random)					
Bozkurt et al.	0.657	0.0220	0.614 to 0.700	32.20					
Richard et al.	0.651	0.0230	0.692 to 0.778	32.20				_	
Selmi et al.	0.755	0.0480	0.661 to 0.849	13.10				-	-
Park et al.	0.732	0.0420	0.650 to 0.814	15.90					
Karsiyakali et al.	0.687	0.0730	0.544 to 0.830	6.60			-		
Total (random effects)	0.709	0.0200	0.670 to 0.748	100.00			•		
					0.5	0.6	0.7	0.8	0.9
Studies including R.I.R.S.	AUC	SE	95% CI	Weight (%, fixed)					
Wang et al.	0.737	0.0480	0.643 to 0.831	14.26			•	_	
Bozkurt et al.	0.690	0.0210	0.649 to 0.731	74.48			-		
Selmi et al.	0.752	0.0540	0.646 to 0.858	11.26					
Total (fixed effects)	0.704	0.0181	0.668 to 0.739	100.00		•			
					0.6	0.7	0.8	3	0.9
Studies including RUS	AUC	SE	95% CI	Weight (%. fixed)					
Erbin et al.	0.655	0.0330	0.590 to 0.720	12.66					
Bozkurt et al.	0.689	0.0210	0.648 to 0.730	31.27					
Richard et al.	0.644	0.0180	0.609 to 0.679	42.56		-	┣─		
Selmi et al.	0.735	0.0520	0.633 to 0.837	5.10		_	-		_
Sfoungaristos et al.	0.707	0.0690	0.572 to 0.842	2.90			-		
Wang et al.	0.700	0.0500	0.602 to 0.798	5.52			-		
Total (fixed effects)	0.669	0.0117	0.646 to 0.692	100.00			•		
					0.5	0.6	0.7	0.8	0.9
Studies including S.T.O.N.E.	AUC	SE	95% CI	Weight (%. fixed)					
Selmi et al.	0.725	0.0500	0.627 to 0.823	22.71	_				
Molina et al.	0.764	0.0320	0.701 to 0.827	55.45		-			
Karsiyakali et al.	0.837	0.0510	0.737 to 0.937	21.83		_			
Total (fixed effects)	0.771	0.0238	0.724 to 0.818	100.00					
					0.6	0.7	0.8	0.9	1.0
Studies including Ito's score	AUC	SE	95% CI	Weight (%. random)					
Richard et al.	0.735	0.0220	0.692 to 0.778	33.42			-	-	
Bozkurt et al.	0.303	0.0200	0.264 to 0.342	33.47	-#-				
Ito et al.	0.870	0.0320	0.807 to 0.933	33.11				_	-
					0.3	0.4 0.5	0.6 0.7	7 0.8	0,9

Table 6. Meta-data extracted from selected studies for all scoring systems and results of meta-analyses

RUS – Resorlu-Unsal score; SReSC – Seoul National University Renal Stone Complexity; AUC – area under curve; SE – standard error; CI – confidence interval

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**Ito's nomogram.** Three studies reported AUCs for Ito's nomogram [4, 19, 22]. Meta-analysis yielded a pooled AUC of 0.635 (95% CI 0.361 to 0.909). Heterogeneity was very high ( $I^2 = 99\%$ ). Statistically significant and very high heterogeneity could not be resolved by exclusion of any study.

# DISCUSSION

There were five scoring systems which had been validated externally and well-studied; S-ReSC, R.I.R.S., S.T.O.N.E. scores, RUS and Ito's nomogram.

S-ReSC score. This scoring system was introduced firstly for the prediction of SFR after percutaneous nephrolithotomy [14]. Then it was modified for outcomes of RIRS by Jung and colleagues [15]. The modified S-ReSC score is calculated sum of 1 points attended for each of following stone locations: the renal pelvis (#1), superior and inferior major calvceal groups (#2-3), and anterior and posterior minor calyceal groups of the superior (#4-5), middle (#6-7), and inferior calyx (#8-9). If the stone is in the inferior sites (#3, #8-9), one additional point per site is added. **R.I.R.S. score.** Xiao and colleagues developed R.I.R.S. score in 2017 [16]. This score is assigned according to following criteria; (R)enal stone density ≤1000 Hounsfield Unit (HU) (1 point) or >1000 HU (2 points), the renal infundibulopelvic angle (RIPA, defined as the inner angle of the intersection of ureteropelvic axis and the axis of the lower renal calyx) of the (I)nferior pole stone (scored from 1 to 3 points as determined by a non-inferior pole stone or inferior pole stone with RIPA >30° or  $\leq$ 30°), (R)enal infundibular length (RIL, the distance from most distal point at bottom stone-containing calix to midpoint of lip of renal pelvis) >25 mm (2 points) or  $\leq$ 25 mm (1 point), (S)tone burden (1 to 3 points for stone burden according to cumulative stone diameter ≤10 mm, >10 mm and  $\leq 20$  mm, and >20 mm, respectively).

**RUS score.** The first developed and validated scoring system for RIRS was introduced by Resorlu and colleagues in 2012 [17]. There are four criteria have equal weight (1 point for each); stone size >20 mm, lower pole stone location and RIPA <45°, stone number in different calyces >1, abnormal renal anatomy (horseshoe kidney or pelvic kidney).

**S.T.O.N.E. score.** This scoring system was derived from pre-operative radiological features of stones [18]. Name of the scoring system is an acronym of included parameters. (S) represents stone size; 1, 2 and 3 points for stones <5 mm, 5-10 mm, and  $\geq 10 \text{ mm}$ , respectively. (T)opography of stone is classified as distal to mid-ureter (1 point), proximal ureter, upper and middle pole of kidney (2 points), lower pole (3 points). (O)bstruction is scored with

the respect of hydronephrosis degree. (N)umber of stone is scored as 1 stone = 1 point, 2 stones = 2 points, >2 stones = 3 points. Finally, (E)valuation of HU is included in the scoring system as follows; <750 HU = 1 point, 750–1000 HU = 2 points and >1000 HU = 3 points.

**Ito's nomogram.** The only nomogram which was included in the systematic review was Ito's. It was introduced by Ito and colleagues in 2014 [19]. Stone volume ( $\leq 500$ ,  $500 < \times \leq 1000$ ,  $1000 < \times \leq 2000$ , >2000 mm<sup>3</sup>), lower pole calculi, operator experience (< 50,  $\geq 50$ ), hydronephrosis, and number of stones are the parameters of nomogram. A total score is calculated according to these parameters (total score 0–25).

This systematic review and meta-analysis has not revealed superiority of any scoring system which aimed to predict surgical outcomes of RIRS. Statistically significant heterogeneity prevented results of analysis from interpreting predictive/discriminative ability of all recent scoring systems. Although heterogeneity was resolved to certain extent by the exclusion of some studies, following evident clinical obstacles will continue to be a source of heterogeneity for further studies; the inter-/intra-observer variability of scores which were calculated by many different scorers, retrospective calculation bias, discordance between studies for the exact diameter of significant residual fragments and surgical outcome assessment methods. Furthermore, some controversial parameters of scoring systems could have contributed the heterogeneity. Stone burden was the most included parameter. Its high predictive performance is still valid but there are controversies about the calculation method of it [27]. Similar to literature, stone burden unit (mm, mm<sup>2</sup>, cm<sup>3</sup>) showed variety among included studies. Also, heterogeneity aside, any scoring system did not demonstrate good predictive/discriminative performance (pooled AUC  $\geq 0.800$ ).

Ozbek's study was a strong source of heterogeneity for three scoring systems; S-ReSC, R.I.R.S. and RUS [20]. AUC values for these scoring systems were reported higher in Ozbek's study than other studies. Xiao's study which was another cause of heterogeneity for R.I.R.S. and RUS scores, was the only article reporting higher AUC values than Ozbek's study [16]. This two studies were responsible a high amount of heterogeneity.

After the exclusion of studies which were source of heterogeneity for each scoring system, S.T.O.N.E. score had higher pooled AUC value (0.771) than S-ReSC, R.I.R.S., and RUS scores (0.709, 0.704 and 0.669, respectively). High heterogeneity could not be resolved by exclusion of any study for Ito's nomogram. Thus, it was not involved further interpretation. Pooled AUC of S.T.O.N.E. score increased and heterogeneity was resolved completely after exclusion of Richard's study [22]. However, an uncertainty arising from Karsiyakali's study compromised the reliability of pooled AUC value. The authors of this study reported high predictive performance in favor of S.T.O.N.E score as a finding of comparative analysis involving four different scoring systems [23]. Although AUC value of S.T.O.N.E score was reported higher than other scores, pairwise analysis did not reveal any superiority for this scoring system in same study. Also the citation which addressed the developers of scoring system was not the developmental study of Molina's. The authors have cited another S.T.O.N.E. score which developed by Okhunov and colleagues for prediction of PNL's surgical outcomes [28]. It was developed at the same time with Molina's system and has very similar design. But Okhunov's system involves some PNL specific parameters such as trach length.

Except R.I.R.S. score, developmental studies of all other scoring systems did not report any comparative findings between their model and recent scoring systems [15–19]. All independent comparative studies emphasized the superiority of a different scoring system via indirect statistical analyses [4, 5, 20, 22]. This is probably due to the difficulties faced in direct comparison analyses. Same difficulty was faced during the further comparative analysis stage of this meta-analysis. At the protocol stage, we decided to perform second stage meta-analyses after obtaining of pooled AUC values for all scoring systems. But these values were derived from mixed patient cohorts (same patients from matched studies and different patients from other studies). Then we tried to have new pooled AUC values derived from same patients for scoring system couples. Also there was another requirement before pairwise comparisons; the independence of scoring systems. All of the scoring systems showed dependency because of the mutual parameters. Only matching of S-ReSC and R.I.R.S. scores provided statistically independency. But further comparison did not require due to quite similar pooled AUC values of these scores. Minimal dependency between Ito's nomogram and S-ReSC may allow direct comparison of them [29].

Retrograde intrarenal surgery may require usage of some instruments during, before or after the procedure. Ureteral access sheath (UAS), stone basket usage or pre/postoperative double j stent placement are the common instrumentations. Their use may influence the outcomes of RIRS [30, 31]. But surgeon attitudes towards this utilizations complicate the surgical technique standardization of RIRS. Most selected studies had reported the UAS/basket usage. But there was no enough metadata about pre/ postoperative double j stent insertion rates. This was another cause for heterogeneity and poor quality beyond all the statistical and reporting problems discussed above.

## **Strengths and limitations**

This is the first systematic review and meta-analysis on comparison of recent scoring systems predicting outcomes of RIRS. We only used AUC metadata from studies to provide inclusive findings. More reliable metadata (for example multivariable analysis results) were not exist in most of the studies. Other limitations were poor quality, the heterogeneity between the studies, cohorts differencies and dependencies between scoring systems preventing further comparisons. All of these limitations affected the generalizability of the findings.

## **Future research**

Despite abundance of scoring system and nomogram studies, ideal system to predict surgical outcomes of RIRS is still an unmet need. A scorer friendly system showing low inter-/intraobserver variability, high discriminative/predictive ability may gain popularity and may be used widely.

# CONCLUSIONS

There was no superiority of any current scoring system which aimed to predict surgical outcomes of RIRS. Although S.T.O.N.E. score showed highest AUC value, high heterogeneity between studies and dependencies between scoring systems make difficult to design a comparative statistical model to generalize these findings. Also, limitations aside, neither scoring system has demonstrated good predictive/ discriminative performance.

### **CONFLICTS OF INTEREST**

The authors declare no conflicts of interest.

#### AUTHOR CONTRIBUTIONS

Research conception and design: OÖ, BÖ, HA. Data acquisition: OÖ, CB, ÖÇ, HMA. Statistical analysis: OÖ, CB. Data analysis and interpretation: HÇ, OÖ, HMA. Drafting of the manuscript: OÖ, EBS. Critical revision of the manuscript: BÖ, HÖ, HÇ, CMY. Obtaining funding: HA, BÖ. Administrative, technical, or material support: HMA, ÖÇ. Supervision: CMY, EBS. Approval of the final manuscript: BÖ, HA.

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